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An Efficient Log-Type Class Of Estimators Using Auxiliary Information Under Double Sampling

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Abstract: In this paper, a class of log-type estimator using the auxiliary information in form of variable is proposed. Double sampling technique has been considered as it is assumed that the auxiliary information about the auxiliary variable is unknown. Bias and mean squared error has been found up to the first order of approximation. The proposed classes are compared to some commonly used estimators both theoretically as well as empirically and they perform better than commonly used estimators available in literature.

Keywords: Mean squared error, efficiency, ratio method of estimation

1 Introduction

Double sampling proves to be a powerful technique when population mean or population total of auxiliary variable is unknown. Moreover in sampling theory, use of auxiliary information, is always beneficial in order to get more efficient estimates of population parameters. Various authors have made the use of auxiliary variable as a source of auxiliary information to increase the precision of the estimators, for the estimation of the population parameter under consideration. In recent years, many authors have also made use of various parameters associated with the auxiliary variable i.e. x for e.g., standard deviation S_x , coefficient of variation C_x , coefficient of kurtosis $b_2(x)$ and correlation coefficient r of the population in estimation of the population variance. [1–8, 10–13] etc are some of the authors in the list. In this paper, a family of estimators have been proposed by adapting the estimator of [15] and a class of log type estimators [14, 16–18] using the auxiliary information on a variable.

Consider a case of finite population $U = (U_1, U_2, ..., U_N)$ of size N from which a large sample of size n' is drawn according to simple random sampling without replacement (SRSWOR) under double sampling. Let \bar{x}' is the larger sample mean which is used to estimate the unknown population mean of auxiliary variable and $s_x^{2'}$ be its sample variance.

Now, another sample of size *n* is taken from *n'* to estimate the sample mean of auxiliary variable only. Let y_i and x_i represent the values of the study and auxiliary variables for the *i*th unit (i = 1, 2, ..., n) of the population. Further, let \bar{y} and \bar{x}_i be the sample means and $s_y^2 = \frac{\sum_{i=1}^n (y_i - \bar{y})^2}{(n-1)}$ and $s_x^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{(n-1)}$ be the sample variance of the study and auxiliary

variables respectively.

2 Proposed estimators

In this paper, the following family of estimators has been proposed for the estimation of the population variance of the study variable *y* using auxiliary information on a variable.

$$T_c = \left[w_1 s_y^2 + w_2 \left(\frac{s_x'^2}{s_x^2} \right) \right] \left[1 + a \log \left(\frac{s_x'^{*2}}{s_x^{*2}} \right) \right] \tag{1}$$

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where *a* is the characterizing scalar.

 $s_x^{*2} = a s_x^2 + b, S_x^{*2} = a S_x^2 + b$

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such that $a_1(\neq 0)$, b_1 and $a_2(\neq 0)$, b_2 are either real numbers or functions of the known parameters of the auxiliary variable x such as standard deviations S_x , coefficient of variation C_x , coefficient of kurtosis b_{2x} , coefficient of skewness b_{1xi} and correlation coefficient r of population.

3 Properties of proposed estimators

In order to obtain the bias and mean square error (MSE), let us consider

 $E(\varepsilon_{0}) = 0 = E(\varepsilon_{1}) = E(\varepsilon_{1}^{'}), E(\varepsilon_{0})^{2} = Ib_{2y}^{*}, E(\varepsilon_{1})^{2} = Ib_{2x}^{*}, E(\varepsilon_{1}^{'})^{2} = I^{'}b_{2x}^{*}, E(\varepsilon_{0}\varepsilon_{1}) = II_{22yx}^{*}, E(\varepsilon_{0}\varepsilon_{1}^{'}) = I^{'}I_{22yx}^{*}, E(\varepsilon_{1}\varepsilon_{1}^{'}) = I^{'}b_{2x}^{*}, E(\varepsilon_{1}\varepsilon_{1}^{'})^{2} = I^{'}b_{2x}^{*}, E(\varepsilon_{1}^{'})^{2} = I^{'}b_{2x$

where, $b_{2x}^* = b_{2x} - 1$, $b_{2y}^* = b_{2y} - 1$ and $I_{22yx}^* = I_{22yx} - 1$; $I_{pq} = m_{pq}/m_{20}^{\frac{p}{2}}m_{02}^{\frac{q}{2}}$, $m_{pq} = \sum_{i=1}^{N} (Y_i - \bar{Y})^p (Y_i - \bar{Y})^q / N$, I = 1/N, I' = 1/n', $b_{2y} = m_{40}/m_{20}^2$, $b_{2x} = m_{04}/m_{02}^2$ are the coefficient of kurtosis of study and auxiliary variable respectively.

Theorem 1.Bias and mean squared error of the proposed estimators are given by

$$\begin{split} Bias(T) &= S_{y}^{2} \left[w_{1} \left\{ 1 + (I - I') \left(a\eta^{2}b_{2x}^{*} - a\eta r_{yx} \sqrt{b_{2y}^{*}b_{2x}^{*}} - \frac{a\eta^{2}}{2} b_{2x}^{*} \right) \right\} - 1 \right] + w_{2} \left\{ 1 + (I - I') \left(b_{2x}^{*} + a\eta Ib_{2x}^{*} + \frac{a\eta^{2}}{2} Ib_{2x}^{*} \right) \right\} \\ MSE(T) &= S_{y}^{4} w_{1}^{2} A + w_{2}^{2} B + S_{y}^{4} w_{1} D + S_{y}^{2} w_{2} G + S_{y}^{2} w_{1} w_{2} F + S_{y}^{4} \\ where \\ A &= 1 + (I - I') \left(b_{2y}^{*} + a^{2} \eta^{2} b_{2x}^{*} - 4a\eta r_{yx} \sqrt{b_{2y}^{*} b_{2x}^{*}} - a\eta^{2} b_{2x}^{*} \right) \\ B &= 1 + (I - I') \left(b_{2x}^{*} + a^{2} \eta^{2} b_{2x}^{*} + 4a\eta r_{yx} b_{2x}^{*} + a\eta^{2} b_{2x}^{*} \right) \\ B &= 1 + (I - I') \left(2a\eta r_{yx} \sqrt{b_{2y}^{*} b_{2x}^{*}} - a\eta^{2} b_{2x}^{*} \right) - 2 \right] \\ G &= \left[(I - I') \left(-a\eta^{2} b_{2x}^{*} - 2a\eta b_{2x}^{*} \right) - 2 \right] \\ F &= 2 + 2(I - I') \left(2a\eta b_{2x}^{*} - 2a\eta r_{yx} \sqrt{b_{2y}^{*} b_{2x}^{*}} - r_{yx} \sqrt{b_{2y}^{*} b_{2x}^{*}} + a\eta^{2} b_{2x}^{*} + a^{2} \eta^{2} b_{2x}^{*} \right) \\ \eta &= \frac{aS_{x}^{2}}{aS_{x}^{2} + b} \end{split}$$

where $\eta_1 = \frac{a_1 S_x^2}{a_1 S_x^2 + b_1}, \ \eta_2 = \frac{a_2 S_f^2}{a_2 S_f^2 + b_2}, \ r_{yx} = \frac{I_{22yx}^*}{\sqrt{b_{2y}^* b_{2x}^*}},$

Corollary 1. *Mean square error of the proposed class of estimator* T_c *is minimum for the optimum value of the characterizing parameters, given*

$$w_{1opt} = \left[\frac{GF - 2BD}{4AB - F^2}\right]$$

$$w_{2opt} = S_y^2 \left[\frac{DF - 2GA}{4AB - F^2}\right]$$
(3)

and the minimum value of the mean square error within the proposed class of estimator is

$$M(T_c)_{opt} = S_y^4 \left[1 - \frac{BD^2 - DFG + AG^2}{4AB - F^2} \right]$$
(4)



4 Some members of the proposed class of estimators

It can be easily seen that a proposed class T_c is a generalized form of class of estimators for constants $a_1 (\neq 0)$, b_1 and $a_2 (\neq 0)$, b_2 is either real numbers or functions of the known parameters of the auxiliary variable x such as the standard deviations S_x , coefficient of variation C_x , coefficient of kurtosis $b_2(x)$, coefficient of skewness $b_1(x)$ and correlation coefficient r of the population. Therefore, a wide variety of estimators can be designed using the above known population parameters. Some of them are listed below.

Table 1: Some members of the proposed class of estimators i.e. T_c					
Log-type estimators T_c	a_1	b_1			
$T_{c_1} = \left[w_1 s_y^2 + w_2 \left(\frac{s_x^{2'}}{s_x^2} \right) \right]^a \left[1 + \log \left(\frac{s_x^{*2'}}{s_x^{*2}} \right) \right]^b$	1	0			
$T_{c_2} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{s_x^{2'} + C_x}{s_x^2 + C_x} \right) \right]^a \left[1 + \log \left(\frac{s_x^{*2'} + C_x}{s_x^{*2} + C_x} \right) \right]^b$	1	<i>C_x</i>			
$T_{c_3} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{b_2(x) s_x^{2'} + C_x}{b_2(x) s_x^2 + C_x} \right) \right]^a \left[1 + \log \left(\frac{b_2(x) s_x^{*2'} + C_x}{b_2(x) s_x^{*2} + C_x} \right) \right]^b$	<i>b</i> _{2<i>x</i>}	C _x			
$T_{c_4} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{C_x s_x^{2'} + b_2(x)}{C_x s_x^2 + b_2(x)} \right) \right]^a \left[1 + log \left(\frac{C_x s_x^{*2'} + b_2(x)}{C_x s_x^{*2} + b_2(x)} \right) \right]^b$	C _x	b_{2x}			
$T_{c_5} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{s_x^{2'} + S_x}{s_x^2 + S_x} \right) \right]^a \left[1 + \log \left(\frac{s_x^{*2'} + S_x}{s_x^{*2} + S_x} \right) \right]^b$	1	S _x			
$T_{6} = s_{y}^{2} \left[w_{1} s_{y}^{2} + w_{2} \left(\frac{b_{1}(x) s_{x}^{2'} + S_{x}}{b_{1}(x) s_{x}^{2} + S_{x}} \right) \right]^{a} \left[1 + \log \left(\frac{b_{1}(x) s_{x}^{*2'} + S_{x}}{b_{1}(x) s_{x}^{*2} + S_{x}} \right) \right]^{b}$	<i>b</i> _{1<i>x</i>}	S _x			
$T_{c_7} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{b_2(x) s_x^{2'} + S_x}{b_2(x) s_x^2 + S_x} \right) \right]^a \left[1 + \log \left(\frac{b_2(x) s_x^{*2'} + S_x}{b_2(x) s_x^{*2} + S_f} \right) \right]^b$	<i>b</i> _{2<i>x</i>}	S _x			
$T_{c_8} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{s_x^{2'} + r}{s_x^2 + r} \right) \right]^a \left[1 + \log \left(\frac{s_x^{*2'} + r}{s_x^{*2} + r} \right) \right]^b$	1	r			
$T_{c_9} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{s_x^{2'} + b_2(x)}{s_x^2 + b_2(x)} \right) \right]^a \left[1 + \log \left(\frac{s_x^{*2'} + b_2(x)}{s_x^{*2} + b_2(x)} \right) \right]^b$	1	$b_2(x)$			
$T_{c_{10}} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{C_x s_x^{2'} + r}{C_x s_x^2 + r} \right) \right]^a \left[1 + \log \left(\frac{C_x s_x^{*2'} + r}{C_x s_x^{*2} + r} \right) \right]^b$	C _x	r			
$T_{c_{11}} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{r s_x^{2'} + C_x}{r s_x^2 + C_x} \right) \right]^a \left[1 + log \left(\frac{r s_x^{*2'} + C_x}{r s_x^{*2} + C_x} \right) \right]^b$	r	C _x			
$T_{c_{12}} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{b_2(x) s_x^{2'} + r}{b_2(x) s_x^2 + r} \right) \right]^a \left[1 + \log \left(\frac{b_2(x) s_x^{*2'} + r}{b_2(x) s_x^{*2} + r} \right) \right]^b$	<i>b</i> _{2<i>x</i>}	r			
$T_{c_{13}} = s_y^2 \left[w_1 s_y^2 + w_2 \left(\frac{r s_x^{2'} + b_2(x)}{r s_x^2 + b_2(x)} \right) \right]^a \left[1 + log \left(\frac{r s_x^{*2'} + b_2(x)}{r s_x^{*2} + b_2(x)} \right) \right]^b$	r	$b_2(x)$			



5 Dominance Condition

Let us now compare the proposed classes of estimators with the conventional estimators.

5.1 Variance estimator in case of SRSWOR

 $MSE(t_1) > MSE(T_c)_{opt}$

5.2 Ratio type variance estimator

 $MSE(t_2) > MSE(T_c)_{opt}$

5.3 Product type variance estimator

 $MSE(t_3) > MSE(T_c)_{opt}$

5.4 Singh et al. (2011) estimator

 $MSE(t_4) > MSE(T_c)_{opt}$

5.5 Isaki (1983) estimator

 $MSE(t_5) > MSE(T_c)_{opt}$

5.6 Das and Tripathi (1978) estimator

 $MSE(t_6) > MSE(T_c)_{opt}$

5.7 Kadilar and Cingi (2006) estimator

 $MSE(t_7) > MSE(T_c)_{opt}$

Thus, it can be easily seen and verified that the proposed class of log-type estimators is far better than these above mentioned estimators which is available in sampling literature.

6 Empirical Performance

The summary and the percent relative efficiency of the following estimators are as follows:



Table 2: Parameters of Data						
Parameters	Population 1	Population 2	Population 3	Population 4		
n'	33	36	22	34		
п	11	11	8	10		
$\begin{array}{c} b_{2y}^* \\ b_{2x}^* \end{array}$	4.032	2.632	9.433	2.725		
b_{2r}^{*}	1.388	2.402	7.105	12.366		
I_{22xy}^{*}	0.305	1.835	8.140	0.224		

 Table 3: PRE of the estimators

Estimator	Population 1	Population 2	Population 3	Population 4	
t_1	100	100	100	100	
t_2	87.167	65.093	162.742	21.544	
<i>t</i> ₃	38.525	13.593	13.009	12.407	
t_4	116.860	248.212	5107.79	67.4228	
t_5	141.940	277.598	574168.1	637.142	
t_6	142.235	293.078	605514.2	666.034	
t_7	142.235	293.078	605514.2	666.034	
T _{iopt}	98.67	185.87	322345.2	456.786	

7 Conclusion

The present study extends the idea regarding the effective use of auxiliary information if the relationship between the study variable and the auxiliary variable is of logarithmic type. The present study provides some novel estimators that are very efficient to some known estimators when such auxiliary information is available. This study is also supported through a numerical study and the results of the study are quite encouraging.

Conflict of Interest

The authors declare that they have no conflict of interest.

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