

# A Comparative Study of Ratio Estimators with Regression Estimator

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**Abstract:** Mean is one of the most important measures of central tendency. There is always an attempt to estimate the population mean precisely and accurately. The ratio and regression estimators are the most frequently used estimators for the population mean. It is well known that the regression estimator always performs better than the classical ratio estimator except when the regression line passes through the origin. If the regression line passes through the origin then both the classical ratio and regression estimators are equally efficient. There has been a continuous effort of improving the ratio estimators. The present paper seeks to explore whether the advanced ratio estimators can outweigh the regression estimator in any sense.

**Keywords:** Ratio estimator, regression estimator, mean squared error, efficiency

## 1 Introduction

In estimation, there is a lot of history. Bakker [1] tried to collect some historical facts where the sample mean were used to predict the population mean. As the statistics advances, the researchers focus to make estimates more accurate and precise. Supplementary information usually increases the efficiency of the estimate. For example, if we want to estimate the production of cereal then the area of cultivation can be used as auxiliary information. It appears that Cochran [2] first time uses the auxiliary variable to estimate the population mean of the study variable by defining the classical ratio estimator. The ratio estimator usually performs well when there is a positive correlation between the study and auxiliary variables.

Shrivastava [3] generalize the ratio estimator of Cochran [2]. Ratio method of estimation is further improved by Walsh [4], Ray and Sahai [5], Sisodia and Dwivedi [6], Bahl and Tuteja [7] and Upadhyaya and Singh [8]. Kadilar and Cingi [9] advances the ratio method of estimation by proposing a new class of estimators. Some other notable works on various kind of ratio method of estimation are Gupta and Shabbir [10], Yan and Tian [11], Subramani and Kumarapandiyan [12], Abid et al. [13], Singh, Vishwakarma and Gangele [14], Sing and Yadav [15], and Tiwari, Bhogal and Kumar [16]. Here we are going to compare some improved classes of ratio estimators with regression estimator to see their validity.

For  $y$  as study variable and  $x$  as an auxiliary variable, some notations that used throughout the article are:

- $N, n$ : population and sample size.
- $\bar{Y}, \bar{X}$ : population means.
- $\bar{y}, \bar{x}$ : sample means.
- $\rho$ : correlation coefficient between  $y$  and  $x$ .
- $S_y^2, S_x^2, S_{yx}$ : population variance, covariance and standard deviation of respective variables.
- $s_y^2, s_x^2, s_{yx}$ : sample variance and covariance.
- $C_y = \frac{S_y}{\bar{Y}}, C_x = \frac{S_x}{\bar{X}}$ : coefficient of variation for  $y$  and  $x$ .
- $\beta_1(x)$ : coefficients of skewness for  $x$ .
- $\beta_2(x)$ : coefficients of kurtosis for  $x$ .

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$$\bullet \lambda = \frac{1-f}{n}, f = \frac{n}{N}, R = \frac{\bar{Y}}{\bar{X}}.$$

## 2 Materials and Methods

Cochran [2] proposed ratio estimator  $t_r = \bar{y}(\frac{\bar{X}}{\bar{x}})$  to estimate the population mean  $\bar{Y}$  when there is positive correlation between the study variable  $y$  and auxiliary variable  $x$ . The mean squared error (MSE) of ratio estimator  $t_r$  is

$$MSE(t_r) = \lambda(S_y^2 + R^2 S_x^2 - 2RS_{yx}) \quad (1)$$

Murthy [17] showed that the ratio estimator  $t_r$  performs better than usual estimator  $t = \bar{y}$  whenever  $\frac{1}{2} \frac{C_y}{C_x} < \rho \leq 1$ .

The linear regression estimator given by Cochran [18] as  $t_{reg} = \bar{y} + b_{yx}(\bar{X} - \bar{x})$ ; where  $b_{yx} = \frac{s_{yx}}{s_x^2}$  is regression coefficient of  $y$  on  $x$ . The MSE of regression estimator  $t_{reg}$  is

$$MSE(t_{reg}) = \lambda S_y^2(1 - \rho^2) \quad (2)$$

If  $\hat{\theta}_1$  and  $\hat{\theta}_2$  are estimators of  $\theta$ , then  $\hat{\theta}_1$  is said to be more efficient than  $\hat{\theta}_2$  if  $MSE(\hat{\theta}_1) < MSE(\hat{\theta}_2)$ .

So, if  $t_r$  is more efficient than  $t_{reg}$  then  $MSE(t_r)$  should be less than  $MSE(t_{reg})$ . From equation (1) and (2) we can find that,

$$MSE(t_r) < MSE(t_{reg}) \text{ if } (RS_x - \rho S_y)^2 < 0 \quad (3)$$

which is not possible i.e. ratio estimator  $t_r$  can never be more efficient than regression estimator  $t_{reg}$ . It can also be concluded from equation (3) that

$$MSE(t_r) = MSE(t_{reg}) \text{ iff } (RS_x - \rho S_y)^2 = 0$$

That is,  $t_r$  is equally efficient to  $t_{reg}$  if and only if  $R = \rho \frac{S_y}{S_x} = \frac{S_{yx}}{S_x^2}$ .

Chakrabarty [19] and Ray & Sahai [5] improved the ratio estimator but their optimal MSEs are the same as regression estimator.

Sisodia & Dwivedi [6] modified the ratio estimator by using coefficient of variation of auxiliary variable by defining  $t_{SD} = \bar{y} \left[ \frac{\bar{X} + C_x}{\bar{x} + C_x} \right]$ . Motivated by Sisodia & Dwivedi [6], Upadhyaya & Singh [8] proposed  $t_{US1} = \bar{y} \left[ \frac{C_x \bar{X} + \beta_2(x)}{C_x \bar{x} + \beta_2(x)} \right]$  &  $t_{US2} = \bar{y} \left[ \frac{\beta_2(x) \bar{X} + C_x}{\beta_2(x) \bar{x} + C_x} \right]$ . Continuing the work, Singh [20] proposed  $t_{S1} = \bar{y} \left[ \frac{\bar{X} + S_x}{\bar{x} + S_x} \right]$ ,  $t_{S2} = \bar{y} \left[ \frac{\beta_1(x) \bar{X} + S_x}{\beta_1(x) \bar{x} + S_x} \right]$  &  $t_{S3} = \bar{y} \left[ \frac{\beta_2(x) \bar{X} + S_x}{\beta_2(x) \bar{x} + S_x} \right]$ , Singh & Tailor [21] defines  $t_{ST} = \bar{y} \left[ \frac{\bar{X} + \rho}{\bar{x} + \rho} \right]$  and Singh et al. [22] as  $t_{Sr} = \bar{y} \left[ \frac{\bar{X} + \beta_1(x)}{\bar{x} + \beta_1(x)} \right]$ . Although the estimator  $t_{SD}$ ,  $t_{US1}$ ,  $t_{US2}$ ,  $t_{S1}$ ,  $t_{S2}$ ,  $t_{S3}$ ,  $t_{ST}$  and  $t_{Sr}$  are more efficient than usual ratio estimator  $t_r$  in certain range but none of them matches with regression estimator  $t_{reg}$ .

Kadilar & Cingi [9] combined the concept of Sisodia & Dwivedi [6], Upadhyaya & Singh [8] and Singh et al. [22] with regression estimator to construct some new type of ratio estimator as

$$t_{KC_i} = [\bar{y} + b_{yx}(\bar{X} - \bar{x})]A_i; i = 1, 2, 3, 4, 5.$$

where  $A_1 = \frac{\bar{X}}{\bar{x}}$ ,  $A_2 = \frac{\bar{X} + C_x}{\bar{x} + C_x}$ ,  $A_3 = \frac{\bar{X} + \beta_2(x)}{\bar{x} + \beta_2(x)}$ ,  $A_4 = \frac{C_x \bar{X} + \beta_2(x)}{C_x \bar{x} + \beta_2(x)}$ ,  $A_5 = \frac{\beta_2(x) \bar{X} + C_x}{\beta_2(x) \bar{x} + C_x}$ .

The MSE of  $t_{KC_i}$  is

$$MSE(t_{KC_i}) = \lambda[R_i^2 S_x^2 + S_y^2(1 - \rho^2)]; i = 1, 2, 3, 4, 5.$$

where  $R_1 = \frac{\bar{Y}}{\bar{X}}$ ,  $R_2 = \frac{\bar{Y}}{\bar{X} + C_x}$ ,  $R_3 = \frac{\bar{Y}}{\bar{X} + \beta_2(x)}$ ,  $R_4 = \frac{C_x \bar{Y}}{C_x \bar{X} + \beta_2(x)}$ ,  $R_5 = \frac{\beta_2(x) \bar{Y}}{\beta_2(x) \bar{X} + C_x}$ .

Kadilar & Cingi [9] showed that  $t_{KC_i}$  perform better than all existing ratio estimators. Let us take the same population used by Kadilar & Cingi [9] to compare the performance of  $t_{KC_i}$  and regression estimator  $t_{reg}$ .

The parameters of population are  $N = 106$ ,  $n = 20$ ,  $\bar{Y} = 2212.59$ ,  $\bar{X} = 27421.70$ ,  $S_y = 11551.53$ ,  $S_x = 57460.61$ ,  $\rho = 0.86$ ,  $C_x = 2.10$ , and  $\beta_2(x) = 34.57$ .

For this population, the calculated values of MSEs are  $MSE(t_{KC_1}) = 2281556$ ,  $MSE(t_{KC_2}) = 2281423$ ,  $MSE(t_{KC_3}) = 2279362$ ,  $MSE(t_{KC_4}) = 2280510$ ,  $MSE(t_{KC_5}) = 2281552$ ,  $MSE(t_r) = 2548180$  and  $MSE(t_{reg}) = 1409557$ .

It is observed that the MSE of  $t_{KC_i}$  are less than that of the usual ratio estimator  $t_r$  but the MSE of regression estimator  $t_{reg}$  is much less than MSEs of  $t_{KC_i}$ .

One can say that  $t_{KC_i}$  improves the ratio estimator but results conclude that  $t_{KC_i}$  actually impair the regression estimator as  $t_{KC_i}$  uses regression estimator i.e. the term  $A_i$  in  $t_{KC_i} = [\bar{y} + b_{yx}(\bar{X} - \bar{x})]A_i$  resulted in increment to MSE of regression estimator.

Similarly, Kadilar & Cingi [23] proposed some other ratio estimators as

$$t_{KC2i} = [\bar{y} + b_{yx}(\bar{X} - \bar{x})]B_i; i = 1, 2, 3, 4, 5.$$

where  $B_1 = \frac{\bar{X} + \rho}{\bar{x} + \rho}$ ,  $B_2 = \frac{C_x \bar{X} + \rho}{C_x \bar{x} + \rho}$ ,  $B_3 = \frac{\rho \bar{X} + C_x}{\rho \bar{x} + C_x}$ ,  $B_4 = \frac{\beta_2(x) \bar{X} + \rho}{\beta_2(x) \bar{x} + \rho}$ ,  $B_5 = \frac{\rho \bar{X} + \beta_2(x)}{\rho \bar{x} + \beta_2(x)}$ .

The MSE of  $t_{KC2i}$  is

$$MSE(t_{KC2i}) = \lambda [P_i^2 S_x^2 + S_y^2 (1 - \rho^2)]; i = 1, 2, 3, 4, 5.$$

where  $P_1 = \frac{\bar{Y}}{\bar{X} + \rho}$ ,  $P_2 = \frac{C_x \bar{Y}}{C_x \bar{X} + \rho}$ ,  $P_3 = \frac{\rho \bar{Y}}{\rho \bar{X} + C_x}$ ,  $P_4 = \frac{\beta_2(x) \bar{Y}}{\beta_2(x) \bar{X} + \rho}$ ,  $P_5 = \frac{\rho \bar{Y}}{\rho \bar{X} + \beta_2(x)}$ .

In numerical illustration  $t_{KC2i}$  perform almost similar to  $t_{KC}$ . One can easily verify that for presented data in Kadilar & Cingi [23], the regression estimator  $t_{reg}$  is far better than  $t_{KC2i}$ .

In a similar manner, many works have been done. Some of them are Yan & Tian [11], Subramani & Kumarapandiyan [12, 24], Abid et al. [13], Abid et al. [25], Abid et al. [26] and Singh & Yadav [15].

### 3 Results and Discussion

We can generalize all the ratio estimators discussed and cited above in two classes as

$$t_{r1} = \bar{y} \left[ \frac{a\bar{X} + b}{a\bar{x} + b} \right] \quad (4)$$

$$t_{r2} = [\bar{y} + b_{yx}(\bar{X} - \bar{x})] \left[ \frac{a\bar{X} + b}{a\bar{x} + b} \right] \quad (5)$$

where  $a(\neq 0)$  and  $b$  are either constant or any auxiliary attribute.

**Theorem 1.** For any value of  $a$  and  $b$ ,  $MSE(t_{r1}) \geq MSE(t_{reg})$ .

*Proof.* The MSE of  $t_{r1}$  can be obtained on the lines of Upadhyaya & Singh [8] as

$$MSE(t_{r1}) = \lambda (S_y^2 + R_{r1}^2 S_x^2 - 2R_{r1} S_{yx}) \quad (6)$$

where  $R_{r1} = \frac{a\bar{Y}}{a\bar{X} + b}$ .

We can see that all the terms in  $MSE(t_{r1})$  are constant except  $R_{r1}$  as  $R_{r1}$  depends on  $a$  and  $b$ .

Differentiate  $MSE(t_{r1})$  with respect to  $R_{r1}$  and equate to zero to get the optimum value of  $R_{r1}$  as  $R_{r1}^o = \frac{S_{yx}}{S_x^2}$ . Put  $R_{r1}^o$  in equation (6), we get the optimum MSE of  $t_{r1}$ .

$$MSE_{opt}(t_{r1}) = \lambda S_y^2 (1 - \rho^2) \quad (7)$$

which is equal to  $MSE(t_{reg})$  as given in equation (2).

Again, assume  $MSE(t_{r1}) < MSE(t_{reg})$ .

$$\text{i.e. } \lambda (S_y^2 + R_{r1}^2 S_x^2 - 2R_{r1} S_{yx}) < \lambda S_y^2 (1 - \rho^2) \quad (8)$$

Simplify equation (8), we get  $(R_{r1} S_x - \rho S_y)^2 < 0$ , which is not possible. Hence,

$$MSE(t_{r1}) \geq MSE(t_{reg}) \quad (9)$$

**Theorem 2.** For any value of  $a$  and  $b$ ,  $MSE(t_{r2}) \geq MSE(t_{reg})$ .

*Proof.* The MSE of  $t_{r2}$  can be obtained on the lines of Kadilar & Cingi [9] as

$$MSE(t_{r2}) = \lambda (S_y^2 + R_{r2}^2 S_x^2 - 2R_{r2} S_{yx}) \quad (10)$$

where  $R_{r2} = \frac{a\bar{Y}}{a\bar{X} + b}$ .

As in Theorem 1, it can be shown that

$$MSE(t_{r2}) \geq MSE(t_{reg}) \quad (11)$$

From Theorem 1 and Theorem 2, it is concluded that regression estimator is more efficient than ratio estimators of both classes. It is worth to mention that, ratio estimators of the classes  $t_{r1}$  and  $t_{r2}$  are biased estimators while regression estimator  $t_{reg}$  is unbiased. Hence, regression estimator is more accurate and precise than ratio estimators of classes  $t_{r1}$  and  $t_{r2}$ . Regression estimator  $t_{reg}$  is one of the simplest estimators of population mean. The results of this comparative study show that the linear regression estimator  $t_{reg}$  always works better than both classes of ratio estimators  $t_{r1}$  and  $t_{r2}$ .

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