# Robustness Study of the Sequential Testing Procedures for the New Weibull-Pareto Distribution 

Surinder Kumar*, Vaidehi Singh and Mayank Vaish<br>Department of Statistics, Babasaheb Bhimrao Ambedkar University, Lucknow, India.

Received: 22 Sep. 2017, Revised: 5 Dec. 2017, Accepted: 22 Jan. 2018
Published online: 1 Mar. 2018


#### Abstract

Sequential testing procedures are developed for testing the hypotheses regarding the parameters of the New Weibull-Pareto Distribution (NWPD). Theoretical expression for the operating characteristics (OC) and average sample number (ASN) functions are derived for the scale parameters of the distribution. The robustness of the SPRT'S in respect of OC and ASN functions is studied, when the distribution under study has undergone a change. The results are presented through Tables and Graphs, so that one can see the numerical evaluated departures in OC and ASN functions.


Keywords: New Weibull-Pareto Distribution, Sequential probability ratio test, Operating characteristics, Average Sample Number, Robustness and Acceptance and rejection region.

## 1 Introduction

Wald (1947), is the first who developed the concept of sequential testing of statistical hypotheses for testing between two simple hypotheses. The concept of sequential testing is heavily dominated by the sequential probability ratio test (SPRT). He derived the theoretical expressions for the operating characteristics (OC) and average sample number (ASN) functions, to study the performance of the SPRT'S.

The SPRT has been applied by various authors, to deal with testing problems, for references, Oakland (1950) developed SPRT for testing the simple vs. simple hypothesis concerning the mean of the negative binomial distribution, Epstein and Sobel (1955) dealt the testing of simple hypothesis problem regarding the mean of one parameter exponential distribution through SPRT, Johnson (1966) applied SPRT for testing the hypothesis for the scale parameter of the weibull distribution when the shape parameter is known, Phatarford (1971) dealt the problem of testing the composite hypothesis for the shape parameter of the gamma distribution through SPRT, when the scale parameter is unknown, Bain and Engelhardt (1982) applied SPRT for testing the hypothesis for the shape parameter of a non-homogenous Poisson process and Chaturvedi et al. (2000) developed SPRT for testing simple and composite hypothesis regarding the parameters of a class of distributions representing various life-testing models. Sevil and Demirhan (2008) developed a group sequential test when response variable has an inverse Gaussian distribution with known parameter.

The robustness of the SPRT in respect of OC and ASN functions has been studied by several authors, when the distribution under consideration has undergone a change, while dealing with various probabilistic models. For references, Harter and Moore (1976) gives sampling plans for reliability tests under the assumption of a constant failure rate and by using Monte Carlo techniques the robustness of the exponential SPRT is studied, when the underlying distribution is a weibull distribution, Montagne and Singpurwalla (1985) investigated the robustness of the sequential life-testing procedure with respect to the risks and the expected sample sizes for the exponential distribution when the life length is not exponential, Hubbard and Allen (1991) applied SPRT on the mean of the negative binomial distribution when the dispersion parameter is known and the robustness of the test to the misspecification of dispersion parameter is studied. Chaturvedi et al. (1998) considered a family of life-testing models and studied the robustness of the SPRT'S for various parameters involved in the model and also generalised the results of Montagne and Singpurwalla (1985).

[^0]
## 2 Set-up of the problem

In this article, we consider the NWPD proposed by Nasiru and Lugnterah (2015) with probability density function (pdf) given by

$$
\begin{equation*}
f(x ; \beta, \partial, \theta)=\frac{\beta \partial}{\theta}\left(\frac{x}{\theta}\right)^{\beta-1} \exp \left\{-\partial\left(\frac{x}{\theta}\right)^{\beta}\right\} ; \quad x>0 \tag{1}
\end{equation*}
$$

and cumulative distribution function (cdf)

$$
\begin{equation*}
F(x ; \beta, \partial, \theta)=1-\exp \left\{-\partial\left(\frac{x}{\theta}\right)^{\beta}\right\} ; \quad x>0 \tag{2}
\end{equation*}
$$

where $\beta$ is a shape and $\partial, \theta$ are the scale parameters, respectively. Weibull and Exponential distributions are the specific cases of (1) for $\partial=1$ and for $\partial=1, \beta=1$, respectively.

In Sections 3, 4, 5 and 6, respectively, we develop SPRT'S for testing the simple null hypotheses for the parameters $\partial$ and $\theta$ involved in the model (1). The robustness of the SPRT'S in respect of OC and ASN functions is studied [see Remarks 3.1, 4.1, 5.1 and 6.1]. In Section 7, the acceptance and rejection regions for $H_{0}$ vs. $H_{1}$ in case of $\theta^{\prime} \theta^{\prime}$ are derived and plotted in Figure 7.1. Finally, in Section 8, the results and findings are presented through Tables and Figures.

## 3 SPRT for testing the hypothesis regarding ${ }^{\prime} \partial^{\prime}$ for known values of $\theta$ and $\beta$

The SPRT for testing the simple null hypothesis $H_{0}: \partial=\partial_{0}$ against the simple alternative $H_{1}: \partial=\partial_{1}\left(\partial_{1}>\partial_{0}\right)$ is defined as

$$
\begin{gather*}
Z_{i}=\ln \left[\frac{f\left(x_{i} ; \beta, \partial_{1}, \theta\right)}{f\left(x_{i} ; \beta, \partial_{0}, \theta\right)}\right]  \tag{3}\\
Z_{i}=\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\left(\partial_{1}-\partial_{0}\right)\left(\frac{x_{i}}{\theta}\right)^{\beta} \tag{4}
\end{gather*}
$$

or,

$$
\begin{equation*}
e^{Z_{i}}=\left(\frac{\partial_{1}}{\partial_{0}}\right) \exp \left\{-\left(\partial_{1}-\partial_{0}\right)\left(\frac{x_{i}}{\theta}\right)^{\beta}\right\} \tag{5}
\end{equation*}
$$

Now, we choose two numbers A and B such that $0<B<1<A$. At the $n^{\text {th }}$ stage, accept $H_{0}$ if $\sum_{i=1}^{n} Z_{i} \leq \ln B$, reject $H_{0}$ if $\sum_{i=1}^{n} Z_{i} \geq \ln A$, otherwise continue sampling by taking the $(n+1)^{\text {th }}$ observation. If $\alpha \in(0,1)$ and $\beta \in(0,1)$ are Type I and Type II errors, respectively, then according to Wald (1947), A and B are approximately given by

$$
\begin{equation*}
A \approx \frac{1-\beta}{\alpha} \text { and } B \approx \frac{\beta}{1-\alpha} \tag{6}
\end{equation*}
$$

The OC function $L(\theta)$ is given by

$$
\begin{equation*}
L(\theta)=\frac{A^{h}-1}{A^{h}-B^{h}} \tag{7}
\end{equation*}
$$

where ' $h$ ' is the non-zero solution of

$$
\begin{equation*}
E\left[e^{Z_{i}}\right]^{h}=1 \tag{8}
\end{equation*}
$$

or,

$$
\begin{equation*}
\int_{0}^{\infty}\left[\frac{f\left(x_{i} ; \beta, \partial_{1}, \theta\right)}{f\left(x_{i} ; \beta, \partial_{0}, \theta\right)}\right]^{h} f\left(x_{i} ; \beta, \partial, \theta\right) d x=1 \tag{9}
\end{equation*}
$$

From (1) and (5), we obtain

$$
\begin{equation*}
E\left[e^{Z_{i}}\right]^{h}=\frac{\partial\left(\frac{\partial_{1}}{\partial_{0}}\right)^{h}}{h\left(\partial_{1}-\partial_{0}\right)+\partial} \tag{10}
\end{equation*}
$$

On substituting (10) in (8), we get

$$
\begin{equation*}
\partial=\frac{h\left(\partial_{1}-\partial_{0}\right)}{\left(\frac{\partial_{1}}{\partial_{0}}\right)^{h}-1} \tag{11}
\end{equation*}
$$

The expression (11) is not very useful for finding the values of OC and ASN functions, hence, we will further evaluate (11) in the following manner to obtain the desired results.

$$
\begin{equation*}
h \ln \left(\frac{\partial_{1}}{\partial_{0}}\right)=\ln \left[1+h\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)\right] \tag{12}
\end{equation*}
$$

Using the expansion of $\ln (1+x),-1<x<1$ in (12), retaining the terms up to third degree in ' $h$ ' and on simplifying, we obtain the real roots of 'h' from (13)

$$
\begin{equation*}
\left\{\frac{1}{3}\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)^{3}\right\} h^{2}-\left\{\frac{1}{2}\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)^{2}\right\} h+\left\{\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)-\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)\right\}=0 \tag{13}
\end{equation*}
$$

The ASN function is approximately given by

$$
\begin{equation*}
E(N \mid \partial)=\frac{L(\partial) \ln B+[1-L(\partial)] \ln A}{E(Z)} \tag{14}
\end{equation*}
$$

provided that $E(Z) \neq 0$, where

$$
\begin{equation*}
E(Z)=\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right) \tag{15}
\end{equation*}
$$

From (14) ASN function under $H_{0}$ and $H_{1}$ are given by

$$
\begin{equation*}
E_{0}(N)=\frac{(1-\alpha) \ln B+\alpha \ln A}{\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)} \tag{16}
\end{equation*}
$$

and

$$
\begin{equation*}
E_{1}(N)=\frac{\beta \ln B+(1-\beta) \ln A}{\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)} \tag{17}
\end{equation*}
$$

Remarks 3.1: Let us consider the problem of testing the simple null hypothesis $H_{0}: \partial=13$ against the simple alternative hypothesis $H_{1}: \partial=15$, for $\alpha=\beta=0.05$. The numerical values of OC and ASN functions are shown in Table 3.1 and their curves are plotted in Figure 3.1(a) and 3.1(b), respectively. It is evident from the Table and Figures that the approximation gives satisfactorily results.

## 4 Robustness of the SPRT for ${ }^{\prime} \partial$ ' when ${ }^{\prime} \theta^{\prime}$ has undergone a change

Let us suppose that the parameter ${ }^{\prime} \theta^{\prime}$ has undergone a change to $\theta^{*}$ and then the probability distribution in (1) becomes $f\left(x ; \beta, \partial, \theta^{*}\right)$. In order to study the robustness of SPRT developed in Section 3 with respect to OC and ASN functions, the values of ' $h^{\prime}$ are obtained by solving the following equation

$$
\begin{gather*}
\int_{0}^{\infty}\left[\frac{f\left(x_{i} ; \beta, \partial_{1}, \theta\right)}{f\left(x_{i} ; \beta, \partial_{0}, \theta\right)}\right]^{h} f\left(x_{i} ; \beta, \partial, \theta^{*}\right) d x=1  \tag{18}\\
\left(\frac{\partial_{1}}{\partial_{0}}\right)^{h} \frac{\beta \partial}{\theta^{*}} \int_{0}^{\infty}\left(\frac{x}{\theta^{\beta}}\right)^{\beta-1} \exp \left[-\left\{\frac{h\left(\partial_{1}-\partial_{0}\right)}{\theta^{\beta}}+\frac{\partial}{\theta^{* \beta}}\right\} x^{\beta}\right] d x=1 \\
\\
\frac{\left(\frac{\partial_{1}}{\partial_{0}}\right)^{h} \partial}{\left\{\frac{h\left(\partial_{1}-\partial_{0}\right)}{\theta^{\beta}}+\frac{\partial}{\theta^{* \beta}}\right\} \theta^{* \beta}}=1
\end{gather*}
$$

$$
\frac{\left(\frac{\partial_{1}}{\partial_{0}}\right)^{h} \partial}{\left(\frac{\theta^{*}}{\theta}\right)^{\beta} h\left(\partial_{1}-\partial_{0}\right)+\partial}=1
$$

Finally, we get

$$
\begin{equation*}
\partial=\frac{(p)^{\beta} h\left(\partial_{1}-\partial_{0}\right)}{\left(\frac{\partial_{1}}{\partial_{0}}\right)^{h}-1} \tag{19}
\end{equation*}
$$

where $p=\frac{\theta^{*}}{\theta}$
The expression (19) is not of much use for calculating the numerical values of OC and ASN functions. In order to handle the situation, we rewrite (19) as

$$
\begin{equation*}
h \ln \left(\frac{\partial_{1}}{\partial_{0}}\right)=\ln \left[1+h\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)\left(\frac{\theta^{*}}{\theta}\right)^{\beta}\right] \tag{20}
\end{equation*}
$$

Using the expansion of $\ln (1+x),-1<x<1$ in (20) and retaining the terms up to third degree in ' $h^{\prime}$ ' and on simplifying, we obtain the following quadratic equation in ' $h$ '

$$
\begin{equation*}
\left\{\frac{p^{3 \beta}}{3}\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)^{3}\right\} h^{2}-\left\{\frac{p^{2 \beta}}{2}\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)^{2}\right\} h+\left\{p^{\beta}\left(\frac{\partial_{1}-\partial_{0}}{\partial}\right)-\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)\right\}=0 \tag{21}
\end{equation*}
$$

where $p=\frac{\theta^{*}}{\theta}$
The Robustness of the SPRT with respect to ASN is studied by replacing the denominator of (14) by

$$
E_{\theta^{*}}(Z)=\int_{0}^{\infty} z f\left(x ; \beta, \partial, \theta^{*}\right) d x
$$

or,

$$
\begin{gathered}
E_{\theta^{*}}(Z)=E\left[\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\left(\partial_{1}-\partial_{0}\right) \frac{x^{\beta}}{\theta^{\beta}}\right] \\
=\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\frac{\left(\partial_{1}-\partial_{0}\right)}{\theta^{\beta}} E\left(x^{\beta}\right) \\
=\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\frac{\left(\partial_{1}-\partial_{0}\right)}{\theta^{\beta}} \frac{\theta^{* \beta}}{\partial} \\
=\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\frac{\left(\partial_{1}-\partial_{0}\right)}{\partial}\left(\frac{\theta^{*}}{\theta}\right)^{\beta} \\
=\ln \left(\frac{\partial_{1}}{\partial_{0}}\right)-\frac{\left(\partial_{1}-\partial_{0}\right)}{\partial}(p)^{\beta}
\end{gathered}
$$

where $p=\frac{\theta^{*}}{\theta}$.
Remarks 4.1: Let us consider the example of testing null hypothesis $H_{0}: \partial=13$ vs. $H_{1}: \partial=15$, for $\alpha=\beta=0.05$. The numerical values of OC and ASN functions are obtained for $p=1, p>1$ and $p<1$, in order to study robustness of the SPRT and are presented in Table 4.1(a) and 4.1(b), respectively. The OC and ASN curves are plotted in Figure 4.1 (a) and 4.1(b), respectively. It follows from Figure 4.1(a) that the OC function curve shifts to left (right) for $p<1(p>1)$ of the curve corresponding to $p=1$ and the similar pattern is followed by the ASN function curve in Figure 4.1(b). It is evident from both the curves that the SPRT is highly sensitive for changes in ${ }^{\prime} \theta^{\prime}$.

## 5 SPRT for testing the hypothesis regarding ${ }^{\prime} \theta^{\prime}$, when ${ }^{\prime} \partial$ ' is known

The SPRT for testing the simple null hypothesis $H_{0}: \theta=\theta_{0}$ against the simple alternative $H_{1}: \theta=\theta_{1}\left(\theta_{1}>\theta_{0}\right)$ is defined as

$$
\begin{equation*}
Z_{i}=\ln \left[\frac{f\left(x_{i} ; \beta, \partial, \theta_{1}\right)}{f\left(x_{i} ; \beta, \partial, \theta_{0}\right)}\right] \tag{22}
\end{equation*}
$$

or,

$$
\begin{equation*}
Z_{i}=\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\partial\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right) x_{i}^{\beta} \tag{23}
\end{equation*}
$$

or,

$$
\begin{equation*}
e^{Z_{i}}=\left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta} \exp \left\{-\partial\left(\frac{1}{\theta_{1}{ }^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right) x_{i}^{\beta}\right\} \tag{24}
\end{equation*}
$$

From (1) and (24), we get

$$
\begin{equation*}
E\left[e^{Z_{i}}\right]^{h}=\frac{\left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta h}}{1+h \theta^{\beta}\left(\frac{1}{\theta_{1}{ }^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)} \tag{25}
\end{equation*}
$$

We get from (25) that

$$
\begin{equation*}
\theta=\left[\frac{\left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta h}-1}{h\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)}\right]^{1 / \beta} \tag{26}
\end{equation*}
$$

The expression (26) is not of much use in calculating the numerical values of OC and ASN functions. Again we may rewrite (26) as

$$
\begin{equation*}
\beta h \ln \left(\frac{\theta_{0}}{\theta_{1}}\right)=\ln \left[1+h \theta^{\beta}\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right)\right] \tag{27}
\end{equation*}
$$

Using the expansion for $\ln (1+x),-1<x<1$ in (27), retaining the terms up to third degree in ' $h^{\prime}$ ' and on simplifying, we obtain the following quadratic equation in ${ }^{\prime} h h^{\prime}$

$$
\begin{equation*}
\left\{\frac{\theta^{3 \beta}}{3}\left(\frac{1}{\theta_{1}{ }^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)^{3}\right\} h^{2}-\left\{\frac{\theta^{2 \beta}}{2}\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)^{2}\right\} h+\left\{\theta^{\beta}\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)-\beta \ln \left(\frac{\theta_{0}}{\theta_{1}}\right)\right\}=0 \tag{28}
\end{equation*}
$$

The ASN function is approximately given by

$$
\begin{equation*}
E(N \mid \theta)=\frac{L(\theta) \ln B+[1-L(\theta)] \ln A}{E(Z)} \tag{29}
\end{equation*}
$$

provided that $E(Z) \neq 0$, where

$$
\begin{equation*}
E(Z)=\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\theta^{\beta}\left(\frac{1}{\theta_{1}{ }^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right) \tag{30}
\end{equation*}
$$

From (29) ASN function under $H_{0}$ and $H_{1}$ are given by

$$
\begin{equation*}
E_{0}(N)=\frac{(1-\alpha) \ln B+\alpha \ln A}{\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\theta^{\beta}\left(\frac{1}{\theta_{1}{ }^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)} \tag{31}
\end{equation*}
$$

and

$$
\begin{equation*}
E_{1}(N)=\frac{\beta \ln B+(1-\beta) \ln A}{\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\theta^{\beta}\left(\frac{1}{\theta_{1}{ }^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)} \tag{32}
\end{equation*}
$$

Remarks 5.1: Let us consider the problem of testing the simple null hypothesis $H_{0}: \theta=12$ against the simple alternative hypothesis $H_{1}: \theta=15$, for $\alpha=\beta=0.05$. The numerical values of OC and ASN functions are shown in Table (5.1) and their curves are plotted in Figure 5.1(a) and 5.1(b), respectively. It is evident from the Table (5.1) and Figures 5.1(a) and 5.1(b) that the approximation gives satisfactorily results.

## 6 Robustness of SPRT for ${ }^{\prime} \theta$ ' when $\partial$ has undergone a change

Let us suppose that the parameter ' $\partial$ ' has undergone a change then the probability distribution in (2.1) becomes $f\left(x ; \beta, \partial^{*}, \theta\right)$. To study the robustness of the SPRT developed in Section 5 with respect to OC and ASN functions, the values of ' $h$ ' are obtained from the following equation

$$
\begin{gather*}
\int_{0}^{\infty}\left[\frac{f\left(x_{i} ; \beta, \partial, \theta_{1}\right)}{f\left(x_{i} ; \beta, \partial, \theta_{0}\right)}\right]^{h} f\left(x_{i} ; \beta, \partial^{*}, \theta\right) d x=1  \tag{33}\\
\left(\frac{\theta_{0}}{\theta_{1}}\right)^{h} \frac{\beta \partial^{*}}{\theta^{\beta}} \int_{0}^{\infty} x^{\beta-1} \exp \left[-\left\{\partial h\left(\frac{1}{\theta_{1}}-\frac{1}{\theta_{0}}\right)+\frac{\partial^{*}}{\theta^{\beta}}\right\} x^{\beta}\right] d x=1 \\
\frac{\left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta h} \frac{\partial}{\theta^{\beta}}}{\left\{\partial h\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right)+\frac{\partial^{*}}{\theta^{\beta}}\right\}}=1 \\
\theta^{\beta}=\frac{\left\{\left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta h}-1\right\}\left(\frac{\partial^{*}}{\partial}\right)}{h\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right)} \\
\theta=\left[\frac{\left\{\left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta h}-1\right\}\left(\frac{\partial^{*}}{\partial}\right)}{h\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right)}\right]^{1 / \beta} \tag{34}
\end{gather*}
$$

The expression (34) is not of much use in calculating the numerical values of OC and ASN functions. Rewrite (34) to obtain the real roots of ' $h$ '.

$$
\begin{equation*}
\beta h \ln \left(\frac{\theta_{0}}{\theta_{1}}\right)=\ln \left[1+h \theta^{\beta}\left(\frac{\partial}{\partial^{*}}\right)\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right)\right] \tag{35}
\end{equation*}
$$

Using the expansion of $\ln (1+x),-1<x<1$ in (35), retaining the terms up to third degree in ' $h^{\prime}$ and on simplifying, we obtain the real roots for ' $h$ ' from the following equation

$$
\begin{equation*}
\left\{\frac{\phi^{3} \theta^{3 \beta}}{3}\left(\frac{1}{\theta_{1}{ }^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right)^{3}\right\} h^{2}-\left\{\frac{\phi^{2} \theta^{2 \beta}}{2}\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)^{2}\right\} h+\left\{\phi \theta^{\beta}\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right)-\beta \ln \left(\frac{\theta_{0}}{\theta_{1}}\right)\right\}=0 \tag{36}
\end{equation*}
$$

where $\phi=\frac{\partial}{\partial *}$
The Robustness of the SPRT with respect to ASN can be studied by replacing the denominator of (29) by

$$
E_{\partial^{*}}(Z)=\int_{0}^{\infty} z f\left(x ; \beta, \partial^{*}, \theta\right) d x
$$

or,

$$
\begin{gathered}
E_{\partial^{*}}(Z)=E\left[\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\partial\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right) x^{\beta}\right] \\
=\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\partial\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right) E\left[x^{\beta}\right] \\
=\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\partial\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right) \frac{\theta^{\beta}}{\partial^{*}}
\end{gathered}
$$

$$
\begin{align*}
& =\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\frac{\partial}{\partial^{*}}\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right) \theta^{\beta} \\
& =\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\phi\left(\frac{1}{\theta_{1}^{\beta}}-\frac{1}{\theta_{0}^{\beta}}\right) \theta^{\beta} \tag{37}
\end{align*}
$$

where $\phi=\frac{\partial}{\partial^{*}}$.
Remarks 6.1: Let us consider the problem of testing null hypothesis $H_{0}: \theta=12$ vs. $H_{1}: \theta=15$ for $\alpha=\beta=0.05$. In order to study the robustness of the SPRT, the numerical values of OC and ASN functions are obtained for $\phi=1, \phi>1$ and $\phi<1$ and are given in Table 6.1(a) and 6.1(b), respectively. The OC and ASN curves are plotted in Figure 6.1(a) and 6.1(b), respectively. It follows from Figure 6.1(a) that the OC function curve shifts to right (left) for $\phi<1(\phi>1)$ of the curve corresponding to $\phi=1$ and the similar pattern is followed by the ASN function curve in Figure 6.1(b). It is evident from both the curves that the SPRT is highly sensitive for changes in ' $\partial$ '.

## 7 Implementation of New Weibull Pareto Distribution (NWPD)

The nature of SPRT in case of NWPD is described as, let $X_{1}, X_{2}, X_{3}, \ldots$ be (iid) random variables from NWPD where $\theta>0$. We wish to test the simple null hypothesis $H_{0}: \theta=\theta_{0}$ vs. simple alternative hypothesis $H_{1}: \theta=\theta_{1}\left(\theta_{1}>\theta_{0}\right)$ having a pre-assigned $\alpha>0, \beta<1$. Let A and B be approximately given by $A \approx \frac{1-\beta}{\alpha}$ and $B \approx \frac{\beta}{1-\alpha}$ and $Z_{i}$ is defined as

$$
Z_{i}=\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)^{\beta}-\partial\left(\frac{1}{\theta_{1}{ }^{\beta}}-\frac{1}{\theta_{0}{ }^{\beta}}\right) x_{i}{ }^{\beta}
$$

where $i=1,2,3 \ldots$ Let $n(\geq 1)$, the SPRT given at (23) can simplify as:
Let us define, $Y(n)=\sum_{i=1}^{n} X_{i}$ and $\mathrm{N}=$ first integer $n(\geq 1)$, for which the inequality $Y(n) \leq c_{1}+d n$ or $Y(n) \geq c_{2}+d n$ holds with the constants

$$
\begin{equation*}
c_{1}=\frac{\ln B}{\partial\left(\frac{1}{\theta_{1}}-\frac{1}{\theta_{0}}\right)}, c_{2}=\frac{\ln A}{\partial\left(\frac{1}{\theta_{1}}-\frac{1}{\theta_{0}}\right)} \text { and } d=\frac{\ln \left(\frac{\theta_{0}}{\theta_{1}}\right)}{\partial\left(\frac{1}{\theta_{1}}-\frac{1}{\theta_{0}}\right)} \tag{38}
\end{equation*}
$$

At the stopping stage, if $Y(N) \leq c_{1}+d N$, we accept $H_{0}$ and if $Y(N) \geq c_{2}+d N$, we reject $H_{0}$ for different values of N , where A and B are the fixed quantities. Figure 7.1 shows the acceptance and rejection region for $H_{0}$ under the case when $H_{0}: \theta=12$ vs. $H_{1}: \theta=15, \partial=2$ and $\alpha=\beta=0.05$. From (38), values of the constants are $c_{1}=0.33316, c_{2}=-0.33316$ and $d=6.694$, respectively. Thus, if $Y(N) \leq 0.33316+6.694 N$, we accept $H_{0}$ and if $Y(N) \geq-0.3316+6.694 N$, we accept $H_{1}$ and at the intermediate stage, we continue sampling.

## 8 Tables and Figures

| Table 3.1: OC and ASN Function for |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\left(H_{0}: \partial=13, H_{1}: \partial=15, \alpha=\beta=0.05\right)$ |  |  |  |  |  |
| $\partial$ | $L(\partial)$ | $E(N)$ | $\partial$ | $L(\partial)$ | $E(N)$ |
| 12.4 | 0.99 | 159.810 | 14.2 | 0.34 | 413.371 |
| 12.6 | 0.99 | 183.473 | 14.4 | 0.22 | 384.705 |
| 12.8 | 0.97 | 212.476 | 14.6 | 0.14 | 346.856 |
| 13.0 | 0.95 | 247.619 | 14.8 | 0.08 | 307.783 |
| 13.2 | 0.91 | 288.857 | 15.0 | 0.05 | 272.012 |
| 13.4 | 0.85 | 334.095 | 15.2 | 0.03 | 241.208 |
| 13.6 | 0.75 | 377.757 | 15.4 | 0.02 | 215.434 |
| 13.8 | 0.63 | 410.695 | 15.6 | 0.01 | 194.095 |
| 14.0 | 0.48 | 423.680 | 15.8 | 0.01 | 176.433 |

S. Kumar et al.: Robustness Study of the sequential testing ...

| Table 4.1(a): OC Function for different values of p <br> $\left(H_{0}: \partial=13, H_{1}: \partial=15, \alpha=\beta=0.05\right)$ <br> $\partial$$p^{2} .96$ |  |  |  |  | $p=.98$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $p=1$ | $p=1.02$ | $p=1.04$ |  |  |  |
| 12.0 | 0.99 | 1.00 | 1.00 | 1.00 | 1.00 |
| 12.2 | 0.98 | 0.99 | 1.00 | 1.00 | 1.00 |
| 12.4 | 0.96 | 0.98 | 0.99 | 1.00 | 1.00 |
| 12.6 | 0.93 | 0.97 | 0.99 | 0.99 | 1.00 |
| 12.8 | 0.87 | 0.94 | 0.97 | 0.99 | 1.00 |
| 13.0 | 0.79 | 0.90 | 0.95 | 0.98 | 0.99 |
| 13.2 | 0.66 | 0.82 | 0.91 | 0.96 | 0.98 |
| 13.4 | 0.51 | 0.71 | 0.85 | 0.93 | 0.97 |
| 13.6 | 0.36 | 0.57 | 0.75 | 0.87 | 0.94 |
| 13.8 | 0.24 | 0.42 | 0.63 | 0.79 | 0.89 |
| 14.0 | 0.15 | 0.29 | 0.48 | 0.68 | 0.82 |
| 14.2 | 0.09 | 0.18 | 0.34 | 0.54 | 0.72 |
| 14.4 | 0.05 | 0.11 | 0.22 | 0.4 | 0.59 |
| 14.6 | 0.03 | 0.06 | 0.14 | 0.27 | 0.45 |
| 14.8 | 0.02 | 0.04 | 0.08 | 0.17 | 0.32 |
| 15.0 | 0.01 | 0.02 | 0.05 | 0.11 | 0.21 |
| 15.2 |  | 0.01 | 0.03 | 0.06 | 0.13 |
| 15.4 |  | 0.01 | 0.02 | 0.04 | 0.08 |
| 15.6 |  |  | 0.01 | 0.02 | 0.05 |
| 15.8 |  |  | 0.01 | 0.01 | 0.03 |
| 16.0 |  |  |  | 0.01 | 0.02 |
| 16.2 |  |  |  |  | 0.01 |


| Table 4.1(b): ASN Function for different values of p <br> $\left(H_{0}: \partial=13, H_{1}: \partial=15, \alpha=\beta=0.05\right)$ <br> $\partial$$p_{0}=.96$ |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $p=.98$ | $p=1$ | $p=1.02$ | $p=1.04$ |  |  |
| 12.0 | 171.036 | 144.504 | 124.632 | 109.377 | 97.377 |
| 12.2 | 198.459 | 165.165 | 140.498 | 121.858 | 107.428 |
| 12.4 | 232.202 | 190.607 | 159.810 | 136.813 | 119.284 |
| 12.6 | 272.750 | 221.868 | 183.473 | 154.911 | 133.412 |
| 12.8 | 318.832 | 259.639 | 212.476 | 176.973 | 150.415 |
| 13.0 | 365.674 | 303.378 | 247.619 | 203.929 | 171.036 |
| 13.2 | 404.026 | 349.827 | 288.857 | 236.622 | 196.137 |
| 13.4 | 422.916 | 391.650 | 334.095 | 275.333 | 226.567 |
| 13.6 | 416.583 | 418.47 | 377.757 | 318.832 | 262.812 |
| 13.8 | 388.984 | 422.174 | 410.695 | 363.045 | 304.262 |
| 14.0 | 350.145 | 402.825 | 423.680 | 400.329 | 348.066 |
| 14.2 | 309.365 | 368.300 | 413.371 | 421.371 | 388.110 |
| 14.4 | 272.012 | 328.311 | 384.705 | 420.285 | 415.719 |
| 14.6 | 240.036 | 289.695 | 346.856 | 398.756 | 423.461 |
| 14.8 | 213.496 | 255.647 | 307.783 | 364.480 | 409.944 |
| 15.0 | 191.702 | 226.935 | 272.012 | 325.954 | 380.629 |
| 15.2 | 173.796 | 203.175 | 241.208 | 288.978 | 343.814 |
| 15.4 | 158.991 | 183.600 | 215.434 | 256.263 | 306.328 |
| 15.6 | 146.638 | 167.418 | 194.095 | 228.481 | 272.012 |
| 15.8 | 136.225 | 153.941 | 176.433 | 205.306 | 242.297 |
| 16.0 | 127.356 | 142.611 | 161.737 | 186.069 | 217.249 |
| 16.2 | 119.725 | 132.990 | 149.409 | 170.060 | 196.355 |
| 16.4 | 113.100 | 124.740 | 138.970 | 156.649 | 178.941 |
| 16.6 | 107.298 | 117.600 | 130.046 | 145.321 | 164.363 |


| Table 5.1: OC and ASN Function for |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\left(H_{0}: \theta=12, H_{1}: \theta=15, \alpha=\beta=0.05\right)$ |  |  |  |  |  |
| $\theta$ | $L(\theta)$ | $E(N)$ | $\theta$ | $L(\theta)$ | $E(N)$ |
| 11.0 | 0.996 | 73.443 | 13.6 | 0.398 | 169.990 |
| 11.2 | 0.994 | 79.719 | 13.8 | 0.311 | 162.483 |
| 11.4 | 0.989 | 86.936 | 14.0 | 0.236 | 152.610 |
| 11.6 | 0.982 | 95.210 | 14.2 | 0.175 | 141.529 |
| 11.8 | 0.970 | 104.619 | 14.4 | 0.127 | 130.189 |
| 12.0 | 0.953 | 115.155 | 14.6 | 0.091 | 119.241 |
| 12.2 | 0.926 | 126.650 | 14.8 | 0.064 | 109.063 |
| 12.4 | 0.888 | 138.693 | 15.0 | 0.045 | 99.822 |
| 12.6 | 0.836 | 150.554 | 15.2 | 0.031 | 91.557 |
| 12.8 | 0.768 | 161.174 | 15.4 | 0.021 | 84.225 |
| 13.0 | 0.686 | 169.303 | 15.6 | 0.013 | 77.747 |
| 13.2 | 0.593 | 173.801 | 15.8 | 0.008 | 72.030 |
| 13.4 | 0.494 | 174.009 |  |  |  |


| Table 6.1(a): OC Function for different values of $\phi$ |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\left(H_{0}: \theta=12, H_{1}: \theta=15, \alpha=\beta=0.05\right)$ |  |  |  |  |  |
| $\theta$ | $\phi=.96$ | $\phi=.98$ | $\phi=1$ | $\phi=1.02$ | $\phi=1.04$ |
| 11.0 | 0.999 | 0.998 | 0.996 | 0.993 | 0.988 |
| 11.2 | 0.998 | 0.997 | 0.994 | 0.989 | 0.980 |
| 11.4 | 0.997 | 0.994 | 0.989 | 0.981 | 0.966 |
| 11.6 | 0.995 | 0.99 | 0.982 | 0.968 | 0.945 |
| 11.8 | 0.991 | 0.984 | 0.970 | 0.949 | 0.914 |
| 12.0 | 0.985 | 0.973 | 0.953 | 0.919 | 0.869 |
| 12.2 | 0.976 | 0.957 | 0.926 | 0.878 | 0.808 |
| 12.4 | 0.962 | 0.933 | 0.888 | 0.821 | 0.731 |
| 12.6 | 0.941 | 0.899 | 0.836 | 0.748 | 0.639 |
| 12.8 | 0.911 | 0.852 | 0.768 | 0.661 | 0.538 |
| 13.0 | 0.869 | 0.790 | 0.686 | 0.563 | 0.436 |
| 13.2 | 0.814 | 0.714 | 0.593 | 0.463 | 0.341 |
| 13.4 | 0.744 | 0.625 | 0.494 | 0.367 | 0.258 |
| 13.6 | 0.661 | 0.530 | 0.398 | 0.282 | 0.191 |
| 13.8 | 0.569 | 0.434 | 0.311 | 0.211 | 0.138 |
| 14.0 | 0.475 | 0.344 | 0.236 | 0.154 | 0.098 |
| 14.2 | 0.384 | 0.266 | 0.175 | 0.111 | 0.068 |
| 14.4 | 0.301 | 0.200 | 0.127 | 0.078 | 0.047 |
| 14.6 | 0.231 | 0.148 | 0.091 | 0.055 | 0.032 |
| 14.8 | 0.173 | 0.107 | 0.064 | 0.037 | 0.021 |
| 15.0 | 0.127 | 0.077 | 0.045 | 0.025 | 0.013 |
| 15.2 | 0.092 | 0.054 | 0.031 | 0.017 | 0.008 |
| 15.4 | 0.066 | 0.038 | 0.021 | 0.010 | 0.005 |
| 15.6 | 0.047 | 0.026 | 0.013 | 0.006 | 0.002 |
| 15.8 | 0.033 | 0.017 | 0.008 | 0.003 | 0.001 |


| Table 6.1(b): ASN Function for different values of $\phi$ <br> $\left(H_{0}: \theta=12, H_{1}: \theta=15, \alpha=\beta=0.05\right)$ <br> $\theta$$\| \phi=.96$ |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| $\phi=.98$ | $\phi=1$ | $\phi=1.02$ | $\phi=1.04$ |  |  |
| 11.0 | 62.344 | 67.479 | 73.443 | 80.396 | 88.503 |
| 11.2 | 66.782 | 72.748 | 79.719 | 87.871 | 97.363 |
| 11.4 | 71.838 | 78.786 | 86.936 | 96.458 | 107.459 |
| 11.6 | 77.614 | 85.713 | 95.210 | 106.231 | 118.743 |
| 11.8 | 84.221 | 93.638 | 104.619 | 117.161 | 130.953 |
| 12.0 | 91.767 | 102.645 | 115.155 | 129.033 | 143.512 |
| 12.2 | 100.337 | 112.746 | 126.65 | 141.351 | 155.452 |
| 12.4 | 109.963 | 123.821 | 138.693 | 153.262 | 165.464 |
| 12.6 | 120.571 | 135.544 | 150.554 | 163.578 | 172.152 |
| 12.8 | 131.917 | 147.307 | 161.174 | 170.968 | 174.460 |
| 13.0 | 143.512 | 158.192 | 169.303 | 174.326 | 172.071 |
| 13.2 | 154.584 | 167.048 | 173.801 | 173.159 | 165.521 |
| 13.4 | 164.106 | 172.734 | 174.009 | 167.777 | 155.953 |
| 13.6 | 170.969 | 174.441 | 169.99 | 159.138 | 144.693 |
| 13.8 | 174.253 | 171.981 | 162.483 | 148.484 | 132.895 |
| 14.0 | 173.528 | 165.837 | 152.610 | 136.981 | 121.378 |
| 14.2 | 168.997 | 156.967 | 141.529 | 125.516 | 110.631 |
| 14.4 | 161.399 | 146.486 | 130.189 | 114.654 | 100.880 |
| 14.6 | 151.754 | 135.389 | 119.241 | 104.691 | 92.183 |
| 14.8 | 141.075 | 124.418 | 109.063 | 95.735 | 84.501 |
| 15.0 | 130.189 | 114.041 | 99.822 | 87.779 | 77.747 |
| 15.2 | 119.666 | 104.506 | 91.557 | 80.755 | 71.816 |
| 15.4 | 109.844 | 95.900 | 84.225 | 74.571 | 66.603 |
| 15.6 | 100.880 | 88.220 | 77.747 | 69.126 | 62.011 |
| 15.8 | 92.815 | 81.405 | 72.030 | 64.323 | 57.960 |




## OC FUNCTION CURVE





## ASN FUNCTION CURVE



Figure: 5.1(b)

## OC FUNCTION CURVE





## References

[1] Bacanli, S. and Demirhan, Y. P. (2008): A group sequential test for the Inverse Gaussian Mean. Statistical Papers, 49, 337-386.
[2] Bain, L. J. and Engelhardt, M. (1982): Sequential Probability Ratio Tests for the shape parameter of NHPP. IEEE Transaction on Reliability, R-31, 79-83.
[3] Chaturvedi, A., Kumar, A. and Surinder, K. (1998): Robustness of the sequential procedures for a family of life-testing models. Metron, 56, 117-137.
[4] Chaturvedi, A., Kumar, A. and Surinder, K. (2000): Sequential testing procedures for a class of distributions representing various life testing models. Statistical papers, 41, 65-84.
[5] Epstein, B. and Sobel, M. (1955): Sequential life tests in exponential case. Ann. Math. Statist., 26, 82-93.
[6] Harter, L. and Moore, A. H. (1976): An evaluation of the exponential and Weibull test plans. IEEE Transaction on Reliability, 25, 100-104.
[7] Hubbard, D. J. and Allen, O. B. (1991): Robustness of the SPRT for a negative binomial to misspecification of the dispersion parameter. Biometrics, 47, 419-427.
[8] Johnson, N. L. (1966): Cumulative sum control chart and the Weibull Distribution. Technometrics, 8, 481-491.
[9] Montagne, E. R. and Singpurwalla, N. D. (1985): Robustness of sequential exponential life-testing procedures. Jour. American Statist. Assoc., 391, 715-719.
[10] Nasiru, S. and Lugnterah, A. (2015): The New Weibull Pareto Distribution. Pakistan Journal of Statistics and Operation Research, 11(1), 101-112.
[11] Oakland, G. B. (1950): An application of sequential analysis to whitefish sampling. Biometrics, 6, 59-67.
[12] Phatarfod, R. M. (1971): A sequential test for Gamma distribution. Jour. American Statist. Assoc., 66, 876-878.
[13] Wald, A. (1947): Sequential Analysis. John Wiley and Sons. New York.


Vaidehi Singh, Deptt. of Statistics, BBAU (A central University), Lucknow- India. She has research experience of 2 years, her research areas are Sequential Analysis, Reliability Theory and Bayesian Inference.


Surinder Kumar, Head, Deptt. of Statistics, BBAU (A central University), LucknowIndia. He is having 22 years research experience in various research fields of Statistics such as Sequential Analysis, Reliability Theory, Business Statistics and Bayesian Inference. Prof. Kumar has published more than 40 research publications in various journals of national and international repute.

Mayank Vaish is Guest Faculty in Geostatistics at Department of Applied Geology, School for Environmental Sciences, Babasaheb Bhimrao Ambedkar University, Lucknow. His research areas are Reliability theory and Sequential testing. He has published 05 Research papers in reputed National and International Journals.


[^0]:    * Corresponding author e-mail: surinderntls@ gmail.com

